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# How to cope with disabilities: Development and psychometric properties of the Coping With Disability Difficulties Scale (CDDS)

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#### How to cope with disabilities: Development and psychometric properties of the

#### **Coping with Disability Difficulties Scale (CDDS)**

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1 2 Abstract 3 4 Purpose/Objective: The aim of this study is to develop and test the psychometric properties of 5 the Coping with Disability Difficulties Scale (CDDS), a scale to measure the coping strategies 6 7 used by people with disabilities to face the disability-related difficulties (both caused by disability itself and by stigma) they encounter in their daily lives. 8 Method/Design: An initial pool of 110 items was developed based on previous literature and 9 the results of a qualitative study using semi-structured interviews. The psychometric 10 characteristics of the CDDS were examined in three samples of people with disabilities (each of 11 which included participants with physical, visual and hearing impairments; total N = 590). 12 13 Results: A final scale of 17 items was obtained. The factor structure of the CDDS was tested 14 and replicated with an adequate fit (*RMSEA* = 0.056; *GFI* = 0.98; *CFI* = 0.98) using confirmatory factor analysis. The internal consistency of the four factors (positive thinking, 15 social sensitization and support, adaptation, and avoidance) were adequate to excellent (with 16 alphas ranging from .68 to .86). 17 **Conclusions/Implications:** To the authors' knowledge, this is the first coping scale that is 18 19 specifically designed for people with disabilities, and it can be highly useful for both research

and applied purposes.

21 Key words: coping, disability, scale, stigma, quality of life

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- 23

24 Impact and implications

25	•	The instrument presented in this paper is, to our knowledge, the first coping scale
26		specifically designed to assess the coping strategies used by people with disabilities.
27	•	Results demonstrate that the Coping with Disability Difficulties Scale (CDDS), shows
28		good psychometric properties to measure the coping strategies used by people with
29		disabilities in the Spanish context.
30	•	The Coping with Disability Difficulties Scale (CDDS) can be used both for research

- The Coping with Disability Difficulties Scale (CDDS) can be used both for research
   and intervention regarding people with disabilities. Our data reveal a factor structure
   that should be tested in different cultural contexts. The CDDS can be used to evaluate
   coping strategies, which allows the adaptation and implementation of training
- 34 intervention programs to promote the more positive coping strategies.

In their daily lives, people with disabilities face numerous issues and barriers that arise, directly or indirectly, from their disabilities. In Spain, a 2017 report by the Spanish National Disability Observatory showed that 35.7% of people who stated that they faced severe limitations in performing daily activities were at risk of social exclusion, while this percentage dropped to 29.7% for people who said they had non-severe limitations and 26.1% for those who said they had no limitations at all (Observatorio Estatal de la Discapacidad, 2017).

Another major problem faced by people with disabilities in Spain is unemployment. According
to a report published by the Spanish National Statistical Institute, in 2016, the unemployment
rate was 28.6% for people with disabilities but 19.5% for people without disabilities (Instituto
Nacional de Estadística, 2017).

Authors such as Fine and Asch (1998) argue that many of the problems and handicaps that
people with disabilities face are not directly caused by their disabilities but rather are the
product of a *disabling environment* that fails to accommodate their needs and builds both
physical and social barriers that lead to exclusion and discrimination (Fine & Asch, 1988). In
Spain, approximately 18% of people with disabilities claim to have been discriminated against
at some point because of their disability (Observatorio Estatal de la Discapacidad, 2017).

However, studies show that differences in the levels of quality of life reported by people with
disabilities tend to be weakly related to their degree of disability or other objective variables.
Most of the variance in quality of life can be attributed to differences in the way that
individuals perceive their situation, themselves, and their opportunities (Albrecht & Devlieger,
1999; Brickman, Coates & Janoff-Bulman, 1978; Sprangers & Schwartz, 1999).

57 The psychological processes that explain why some individuals are better off emotionally than 58 others afflicted by the same disability have been described in the literature as "coping". Richard 59 Lazarus and Folkman (1984) defined coping as the "constantly changing cognitive and 60 behavioral efforts to manage specific external and /or internal demands that are appraised as 61 taxing or exceeding the resources of the person". More recently, coping has been described as

"an organizational construct used to encompass the myriad actions individuals use to deal with
stressful experiences" (Skinner, Edge, Altman, & Sherwood, 2003). Thus, coping is not a
unidimensional behavior (Pearlin & Schooler, 1978).

As stated in the previous section, many of the barriers encountered by people with disabilities 65 are a product of social stigma. Social stigma has been identified as a source of stress, and some 66 authors argue that people cope with this stress by using various strategies to protect or enhance 67 the personal and social aspects of their identity (Berjot & Gillet, 2011). For instance, people 68 may attribute their negative experiences to discrimination to protect their personal identity or 69 decrease the importance of a given social identity to protect the social aspects of their identity, 70 71 or they could try to re-evaluate a threatened dimension or affirm their personal qualities to 72 enhance the social and personal aspects of their identity, respectively (Berjot & Gillet, 2011).

#### 73 Coping scales and disability

Previous research has adopted different approaches in the study of coping among people with
disabilities. Some studies have exclusively used qualitative methodologies (Persson, Lars-

76 Oloff, & Ryden, 2006; Boerner, & Wang, 2012) to explore coping strategies.

Some instruments, such as the Freiburg Questionnaire of Coping with Illness (FQCI, Muthny,
1989), take a quantitative approach in studying coping among patients with physical disability
(Haase, Linenemann, & Faustmann, 2008).

80 Other research has employed general coping scales. Kara and Açıkel (2011), in a sample of

patients with physical disabilities, used the Multidimensional Scale of Perceived Social Support
and the Coping Strategy Indicator (Amirkhan, 1990).

83 One of the most widely employed questionnaires has been the Coping Strategies Questionnaire

84 (CSQ; Rosenstiel Keefe, 1983), which has been used in different ways and for people with

85 different kinds of physical disability. Some authors have validated a shortened version of the

original scale, the CSQ 24 (Harland & Georgieff, 2010). Another study (Regier and Parmelee,

2015) combined one subscale of the CSQ that had been shown to be relevant to coping in a

sample of patients with physical disability, the Coping Self-Statements (CSS), with the Coping
with Illness Scale (CWI) (Felton & Revenson, 1984) to create a final version composed of 57
items.

91 The brief version of COPE (Carver, 1997) has also been used in research (Pande, & Tewari,
92 2011; Yuan, Zhang, & Li, 2017). Finally, studies have explored coping by combining
93 qualitative and quantitative approaches (Pande, & Tewari, 2011; Senthil et al., 2018).

#### 94 Goals of this study

95 As stated before, the literature shows that coping is not a unitary construct but rather an 96 organizational construct used to define several different behaviors that people use to deal with 97 stressful experiences. Thus, there is no agreement in the classification of the many possible 98 coping strategies or a universally accepted method of assessment, and there is no consensus on 99 which types of coping behaviors are most effective and under which circumstances.

100 Therefore, to ensure the relevance of the models and measures of coping, they should be based 101 on the specific populations and tailored for those contexts in which they will be applied (Somerfield, 1997). For this purpose, it is necessary to gather first-hand information about how 102 people face their disability, as even caregivers' view on what constitutes good coping may not 103 correspond to that of care receivers (Slöteen, Kreuter, Lampic, & Persson 2005). Different 104 105 coping strategies are identified in each context and in connection with different problem areas. Nonetheless, as described above, studies about coping and disability have used either general 106 107 coping scales or scales that were developed for other areas, such as illness.

The aim of this study was to develop and validate a scale to measure the coping strategies used
by people with disabilities (physical, visual and hearing impairment) to face daily difficulties
caused by the disability itself as well as social barriers caused by stigma (Persson, Lars-Olof &
Ryden, 2006). Because coping strategies are context specific, items were drawn not only from
the previous literature but also from qualitative interviews conducted with a sample of people
with disabilities.

# 115

#### Methods

#### **Participants** 116

117 For the qualitative study, participants were recruited through several organizations for people 118 with disabilities to participate in semi-structured in-depth interviews.

The sample in the pilot study was composed of UNED (National Distance Education 119

120 University) students recruited online through a link at UNIDIS, the assistance service for

121 students with disabilities at the university.

122 Finally, the sample for the correlational ex post facto study was recruited through two different sources. Some participants were recruited from the same organizations that were contacted for 123 124 the qualitative study. The rest of the sample was recruited through undergraduate students of two different courses at UNED, who were asked to send the link with the questionnaire to 125 people with disabilities in exchange of extra course credit. No differences in any of the coping 126 127 scores were found between participants recruited from the two sources.

Regardless of the recruitment method and type of disability, participants were required to be 128 129 over 18 years old and have a certified disability degree of 33% or greater (according to the 130 Spanish administration), which grants them a Disability Certificate, giving them access to 131 certain benefits, rights and services. Participants who did not meet these criteria were excluded 132 from the sample. All studies were conducted in Spanish, and all participants were Spanish speakers living in Spain. As convenience sampling method was used, we tried to obtain a 133 134 sample as large as possible, to alleviate the possible lack of representativeness derived from not 135 using probabilistic sampling. Items were translated into English for publication purposes, and then translated back into Spanish. Both Spanish versions were assessed, and no significant 136 change in meaning was found. 137

138 Table 1 summarizes the sociodemographic and clinical characteristics of the participants in 139 each sample. As the table shows, only people with either physical, hearing, or visual

140 impairments were interviewed in the qualitative study, but people with other types of

141 disabilities were not excluded from the sample in the quantitative studies.

#### 142 INSERT TABLE 1

#### 143 **Procedure**

This research took place from December 2014 to February 2017. All research procedures were
approved by UNED's Ethics Committee. Participants in all phases read a plain language
statement and signed an informed consent form before proceeding with the research. The
Coping with Disability Difficulties Scale (CDDS) battery was developed through the following
steps (Eignor, 2001; Revicki, et al., 2007).

149 After a review of the literature on coping and disability, a qualitative study using the in-depth interview technique was carried out to learn about the perspectives of people with disabilities 150 on the difficulties arising from their condition and how they coped with these difficulties. A 151 total of 27 individual interviews with lengths of 30 to 60 minutes were conducted. These 152 153 interviews used a semi-structured format, which means the interviewer allowed participants to speak freely and in their own words, while ensuring all the relevant topics were covered 154 (Minichiello, Aroni, Timewell, & Alexander; 1990). Then, based on the contents of the 155 156 interviews and the reviewed literature on coping, (1) five members of the research team defined 157 the constructs (the coping strategies used by people with disabilities) after deliberation (Lord & Novick, 1968; Nunnally & Bernstein, 1994). (2) Items were developed to measure each 158 159 construct, following psychometric recommendations (Haladyna, Downing, & Rodriguez, 2002; 160 Osterlind, 1989). (3) A decision on the response format of the items was made. (4) Each item was rated based on its clarity, representativeness and relevance. (5) Based on the experts' 161 162 responses and comments, items were selected and reworded to compile the initial item pool. Then, a cognitive debriefing study was carried out, in which eight people with different types 163 of disabilities rated the questionnaire's understandability and accessibility. 164

Third, a pilot study was conducted to assess the initial items' psychometric properties and 165 166 select those that would be part of the final scale. This study served to purge the item pool based 167 on empirical criteria (e.g., missing values, floor or ceiling effects, internal consistency, or corrected homogeneity index). Participants completed an online questionnaire through 168 Qualtrics. An adapted version of the questionnaire was developed for participants with visual 169 170 impairments using screen readers, but its contents were identical to those of the general version. Finally, to assess the factor structure and to obtain diverse evidence of validity, a correlational 171 ex post facto study was performed. The variables were recoded to eliminate participants with 172 any missing data, which confirmatory factor analysis (CFA) modeling programs cannot 173 accommodate with ease. The final sample analyzed consisted of 520 participants. 174 Participants were randomly divided into two subsamples: First, an exploratory factor analysis 175 176 (EFA) was adopted for subsample 1 (N = 260), and second, we tested the validity of the factor structure derived from the EFA results with subsample 1 using CFA on subsample 2 (N = 260) 177 within the framework of structural equation modeling. Evidence of convergent validity was 178 assessed in the total sample. As in the pilot study, this cross-validation study was conducted 179 180 online using Qualtrics.Measures In the qualitative study, in-depth semi-structured interviews were conducted. Three questions 181 addressed coping strategies specifically: a) "How have you adapted to [your disability]?" b) 182 "To what extent do you feel able to cope with your problems? How do you do it?" and "c) "Is 183 there something that makes you feel better or makes it easier for you to face this problem?" 184 185 The online questionnaire used for the pilot study included the 46 items selected after the interrater process and a set of items that collected information about participants' disability (type, 186 severity and onset) and sociodemographic data (gender, age, educational level, marital status 187 and employment status). The instructions, for all versions of the scale read: "Please, mark how 188 often you use each of the following strategies to cope with difficulties derived from your 189 disability". 190

191 The questionnaire for the two validation studies included the same sociodemographic items,

along with the 24-item version of the CDDS obtained after the psychometric depuration and a

selection based on substantive criteria. Moreover, this questionnaire includes the followingscales to assess convergent and discriminant validity:

*Multidimensional Perceived Discrimination Scale* (Molero, Recio, García-Ael, Fuster &
Sanjuan, 2013). This 16-item scale measures the perception of four different types of
discrimination: blatant group discrimination, subtle group discrimination, blatant personal
discrimination, and subtle personal discrimination. We grouped the four factors into one, which
better served the purpose of our study. In our study, the reliability of the scale was found to be
high (Cronbach's alpha = .94).

*The Stigma Scale for Chronic Illness 9-Item Version (SSCI-9).* We used the Spanish adaptation
of the Internalized Stigma subscale of the SSCI (Rao et al., 2009), which showed good
psychometric properties in people with different types of disabilities (Silván-Ferrero, Recio, &
Nouvillas-Pallejà, 2018). The SSCI-9 was administered with the following response format to
assess frequency: 1 = Never or almost never, 2 = Rarely, 3 = Many times, and 4 = Always or
almost always. In our study, the scale showed good internal consistency (Cronbach's alpha =
.90).

*Group Identification.* We measured identification using a previously validated six-item
measure (Ashforth & Mael, 1989; Mael & Ashforth, 1992). Participants indicated their
agreement with each item on a 4-point Likert scale. In our sample, the internal consistency of
the scale was good (alpha = .82).

Activism. This was measured with four items assessing the perceived effectiveness of collective
action and the intention to engage in it (Pérez-Garín et al., 2017). Participants responded on a
4-point Likert scale, ranging from 1 (totally disagree) to 4 (totally agree). The coefficient alpha
for this scale was .82.

216 Social Support. The Spanish adaptation (Bellón, Delgado, Luna & Lardelli, 1996) of the Duke-

217 UNC Functional Social Support Questionnaire (Broadhead, Gehlbach, Degruy & Kaplan,

218 1988) was used to measure perceived social support. This 11-item scale showed high internal
219 consistency (alpha = .91).

Self-esteem. This construct was measured using the Spanish adaptation (Expósito & Moya,
1999) of the Rosenberg Self-esteem Scale (Rosenberg, 1965). This widely used self-esteem
measure and is composed of 10 items addressing a person's sense of worth and personal value.
Responses are given on a Likert scale from 1 (totally disagree) to 4 (totally agree). Cronbach's
alpha was .85 in the present sample.

225 *Resilience*. This was evaluated using the Connor-Davidson Resilience Scale, 10-item version

226 (Connor & Davidson, 2003), a self-administered questionnaire. Responses are provided on a

227 Likert-type additive scale (from 0 = never to 4 = almost always), which had a single dimension

in the original version. We used the Spanish version (Notario-Pacheco et al., 2011), which

showed adequate reliability in our sample (alpha = .90).

#### 230 Data analyses

Qualitative data were analysed using the constant comparison technique based on Grounded
Theory (Glaser & Strauss, 1967). Core categories were taken out by descriptive coding analysis
using the software Atlas Ti.7. Six experts (in qualitative research and research on disability; all
of whom are part of the research team and colleagues within the same faculty, and one of
whom has experience in community work with stigmatized groups) assessed the
methodological and theoretical quality of core categories and subcategories (Przyborski &
Wohlrab-Sahr, 2010).

In the pilot study carried out for the depuration of the original item pool, the basic

239 psychometric properties of the instruments were assessed using SPSS 24.

In the correlational ex post facto study carried out to obtain evidence of construct and 240 241 convergent validity, the following steps were followed. First, two random subsamples were 242 extracted. Then, to explore the initial structure of the instrument, an EFA using unweighted least squares with promin rotation methods was conducted on subsample 1 (N = 260). A 243 polychoric correlation matrix was used because it is particularly suitable for items with a 244 245 Likert-type response format (Muthen & Kaplan, 1992). The Factor 10.3 program (Lorenzo-Seva & Ferrando, 2006) was used for this purpose. A cut-off for saturation of  $\geq$  .40 was 246 247 established for selecting the items.

248 Then, a CFA was conducted in subsample 2 (N = 260) using the unweighted least squares extraction method (ULS).PRELIS and LISREL 8.7 software were used. To assess the fit of the 249 250 models to the data, several absolute and relative fit indexes are reported (Bentler & Hu, 1995), including chi-square ( $\gamma^2$ ), degrees of freedom, incremental fit indexes (CFI, IFI, NFI, GFI and 251 AGFI), and two residual fit indexes (RMR and RMSEA). Second, to obtain convergent 252 validity, the second sample was analyzed by calculating the Pearson correlation of each factor 253 in the CDDS with quality of life, perceived discrimination, internalized stigma, group 254 identification, activism, social support, self-esteem, and resilience. Approach strategies were 255 256 expected to be negatively correlated with internalized stigma and positively correlated with quality of life, group identification, activism, social support, self-esteem and resilience. 257 258 Avoidance strategies, on the other hand, were expected to be negatively correlated with the latter variables and positively correlated with perceived discrimination and internalized stigma. 259 To compare alternative models we also reported the model CAIC and the significance of the 260 change in Chi-Square  $\Delta \chi^2$  ( $\Delta df$ ). The CAIC is a good fit index for model comparison and the 261 general rule is that values smaller the better (Schreiber, Stage, King, Nora, & Barlow, 2006). 262 Significantly changes in Chi-Square indicate there are differences in the fit of the models 263 compared. 264

Finally, item response theory analysis was used to examine psychometric properties of the 265 CDDS. IRT is a useful tool for gaining insights that traditional techniques cannot provide. It is 266 267 a model-based measurement theory that aims to show the relationship between responses to items and the ability or trait that each item is supposed to be measuring (Embretson v Reise, 268 2010). For a trait, CTT provides a single score, which is derived from the scores of the different 269 270 items comprising the scale, while IRT provides trait scores at the item level. So, the IRT approach would allow the identification of items that are functioning differently in terms of 271 their ability to discriminate and also represent the underlying trait and measure it at different 272 273 levels (Vahedi, 2010). This, in turn, can facilitate the development and revision of the 274 measures.

275 We use the SGR model, appropriate for ordered response categories scales data using IRTPRO 276 4.2 software for the IRT analysis. After having confirmed that the pertinent assumptions were met to perform parametric tests on the definite sample (except normal multivariate kurtosis) we 277 278 used t test and MANOVA to study the differences in CDDS scores according to participant characteristics. Although the subscales did not exhibit normal multivariate kurtosis, according 279 to (Finch, 2005) the parametric statistic is robust even when the assumption of normality is 280 281 violated, and it slightly outperforms the nonparametric statistic in terms of type I error rate and 282 power.

#### 283 **Results**

#### 284 Development of the CDDS items

285 The initial item pool was developed through the following steps:

(a) Definition of the constructs. A theoretical conceptualization of the various coping
strategies analyzed was carried out taking into account the content analysis and the
literature review. Five members of the research team wrote independent definitions and

then met to reach an agreement about the definition of each strategy. A total of 23 289 290 coping strategies were developed.

291 (b) Development and writing of the items: The same five researchers who defined the constructs independently wrote three to four items for each coping strategy and met 292 again to agree on the number of items of the initial battery. The goal of this process was 293 294 to have a representative sample of the universe of possible coping strategies. Psychometric recommendations for the development of items were followed. Criteria of 295 296 relevance (i.e., the content should be clearly related to the construct), clarity (i.e., the 297 items should be written to avoid excessive generality and be written in short, simple and intelligible sentences), and representativeness (i.e., the items should be representative of 298 299 the construct) were taken into account. This process produced an initial pool of 110 items, which were sent to the rest of the team for evaluation. 300 301 (c) Expert assessment and inter-rater process: Four members of the research team evaluated 302 the items in terms of their relevance, clarity and representativeness and suggested

303 rephrasing when necessary. Additionally, each rater chose the best items for each 304 strategy (a maximum of three) in order to develop a short scale to avoid overburdening 305 respondents. After this process, a total of 46 items representing the 23 coping strategies 306 were selected. The entire research team agreed on using a four-point response format

("never or almost never", "rarely", "fairly often", "always or almost always"). 307

(d) Cognitive debriefing interview: Eight people with different types of disabilities 308

309 answered the questionnaire, and they later informed a member of the research team 310 about possible difficulties in understanding the questionnaire and the accessibility of the questionnaire's format. The items were generally regarded as relevant, accessible, and

easy to understand and answer. 312

311

313 Pilot study to analyze the psychometric properties of the items

Eighty-four people with different types of disabilities answered the preliminary version of the CDDS. Four participants were excluded from the analyses for having missing values in three or more items in the questionnaire.

First, items that demonstrated a ceiling or a floor effect were removed. If more than 25% of participants answered with either the lowest or highest point on the scale, the removal of the item was considered.

Because more than one item per strategy had been developed and the goal was to obtain a short battery that was accessible for people with disabilities, both internal consistency and substantive criteria were used to select the final battery of items. After this process, 20 items were selected, representing 17 coping strategies.

324

#### 4 Construct and criterion validity of CDDS

#### 325 Construct validity: Results of the EFA

The factor structure of the remaining items was assessed through EFA in one of the 326 random subsamples obtained (Table 2). A four-factor structure was found, which explained 327 328 55% of the variance. The first factor (F1) included fundamentally cognitive strategies related to positive thinking and acceptance of the situation, and the second factor (F2) grouped social 329 sensitization and social support strategies, vindication of rights and social support. The third 330 331 factor (F3) comprised strategies to adapt to daily life. The fourth factor (F4) comprised avoidance strategies. A graphic representation of the model can be seen in Figure 2. The model 332 fit indexes were Barlett = 1982.5, df= 190, p < .001; KMO = .906, GFI = .99, and CFI: .95, 333 which indicated a very good fit. Three items did not reach the cut-off saturation criteria 334 established for selection ( $\geq$  .40); thus, they were removed. 335

336 INSERT TABLE 2

#### 337 INSERT FIGURE 2

338 Construct validity: Results of the CFA

- To assess the fit of the CDDS to the factor structure obtained in the EFA, CFA was conducted in the second random subsample (n = 260). The results confirmed the four-factor structure with a good fit to the data: Chatorra-Bentler  $\chi^2 = 206.18$  (113), p < .001; *RMSEA* (IC 90%) = 0.056 (0.044; 0.068); *SRMR* = 0.061; *GFI* = 0.98; *AGFI* = 0.97; *CFI* = 0.98; *NFI* = 0.97; *IFI* = .98. Furthermore, all standardized loadings were greater than 0.5, a level considered acceptable (Green & Carroll, 1978). The results of the fully standardized solution of the model
- are displayed in Table . Covariances among the factors are shown in Table 3.

#### 346 INSERT TABLE 3

We compared three alternative models to the four-factor proposed model. As can be seen in
Table 5, there are significant differences in chi-square between the three alternative models and
the proposed model. Furthermore, the alternative models showed a poorer fit than the proposed
one.

#### 351 INSERT TABLE 4

352 Internal consistencyTable 5 shows that the factors with the highest internal consistency were positive thinking and social sensitization and support. The avoidance and adaptation factors had 353 Cronbach's  $\alpha$  values below .70, though these values were acceptable considering the small 354 number of items comprising each factor. Anyway, Jöreskog rho (omega) is a better reliability 355 356 measure than Cronbach's alpha in Structural Equation Modeling, since it is based on the loadings rather than the correlations observed between the observed variables. 357 358 Estimates of the reliability of the four subscales were slightly higher using omega instead of the alpha coefficient, as the alpha statistic underestimates reliability in ordinal data (Bentler, 2009). 359 Given that the threshold of acceptability for omega reliability is .70, the results were 360 361 satisfactory, ranging from .76 through .90 for all the four factors.INSERT TABLE 5

362 *Convergent validity* 

Positive correlations were found for the factors positive thinking, social sensitization and support, and adaptation with quality of life, group identification, activism, social support, selfesteem and resilience, while their relationships with internalized stigma were negative. The avoidance factor, on the other hand, showed the opposite pattern (except for group identification, for which no significant correlation was found) and was also positively related to perceived discrimination (Table 3).

369 CCDS Scores

Descriptive results of the factors in CDDS are shown in Table 3. Positive thinking had the
highest score, while avoidance had the lowest, although it was near the theoretical mean of the
scale. The scores for social sensitization and support as well as adaptation were above their
theoretical means.

#### 374 Differences in CDDS scores according to participant characteristics

Several differences in the CDDS factor scores were found according to participants' type of disability. Participants with visual impairment scored higher on the three factors related to approach strategies (positive thinking, social sensitization and support, and adaptation: F =4,35, p < .01; F = 4.78, p < .01; and F = 5.88, p < .01, respectively). No significant differences were found in the avoidance factor.

380 Women scored significantly higher than men in the social sensitization and support and

adaptation factors (t = 3.37, p < .01 and t = 2.19, p = .03, respectively), and marginal gender

- differences were found in avoidance (t = -1.89, p = .06).No significant gender differences were found for positive thinking.
- 384 Differences were also found in all factors according to education level. The higher the level of
- education was, the higher the scores in the approach factors (positive thinking, social

sensitization and support, and adaptation: F = 2.82, p = .04; F = 3.90, p = .01; and F = 5.60, p < .01

.01, respectively) and the lower the score in avoidance strategies (F = 3.29, p = .02).

People who actively participated in non-governmental organizations (NGOs) in the disabilities field (N = 200) scored significantly higher in the approach factors (positive thinking, social sensitization and support, and adaptation: t = 3.05, p < .01; t = 8.96, p < .01; and t = 2.67, p <.01, respectively) and lower in the avoidance factor (t = -2.97, p < .01).

Significant differences were also found regarding the acquisition of disabilities. Participants who had a disability at birth scored higher than participants with an acquired disability in all three approach factors (t = 2.20, p = .03; t = 4.40, p < .01; t = 2.46, p = .01) and scored lower in the avoidance factor (t = -2.33, p = .02).

#### 396 IRT Analyses

We use the SGR model, appropriate for ordered response categories scales. For this model, the probability of endorsing a response option depends on the discriminating power of the item and the location of the threshold parameter for that option on the latent trait continuum. For a fouroption measure, the SGR model estimates a unique slope parameter for each item and 4-1 = 3threshold parameters. Each threshold reflects the level of general coping strategies needed to have equal (.50) probability of choosing to respond above a given threshold.

We use a unidimensional model because in its strictest sense unidimensionality is usually not 403 404 necessary to take advantage of the benefits of IRT, provided a dominant first dimension is 405 present (Smith & Reise, 1998) as appears to be the case for the current data. As shown in Table 4, results from a one-factor CFA model found a reasonable fit to the sample data,  $(\gamma^2/df =$ 406 175.34 (6); CFI = .96, NFI = .94, SRMR = .086) although it was not the best possible fit to the 407 data. Essential unidimensionality was confirmed using EFA, in which the ratio of eigenvalues 408 of the first and second factors was approximately 6.2:1.5 and the variance accounted for by the 409 410 dominant factor was 36%.

411	The SGR model	was fit to each	item on the	CDDS scale (	(see Table 6	5). Slope	parameters ra	anged
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412 from .62 (item 12) to 3.10 (item 16). Threshold parameters ranged from -4.20 to -1.15 for  $b_1$ , -

413 1.90 to 0.52 for  $b_2$ , and -0.23 to 3.87 for  $b_3$ .

#### 414 INSERT TABLE 6

415	We use the Pearson's Chi-square $(S-\chi^2)$ fit statistics (Orlando & Thissen, 2000) to assess the
416	degree of similarity between predicted and observed response frequencies for each item
417	response category (see Table 6). Results showed a satisfactory fit in that only 4 of the 17 items
418	showed misfit using the .01 criteria (items 7, 10, 16 and 20).
419	As Figure 3 shows, the response category system is operating as expected for each item,
420	because each increasing category is more likely to be selected than previous response
421	categories as one moves along the latent trait continuum.
422	INSERT FIGURE 3
423	Figure 4refers to the total information curve, which is represented by the sum of the
424	information from all the items. It demonstrates that this scale presents greater precision in low
425	and intermediate levels of the latent trait in order to measure coping strategies used by people
426	with disabilities. From trait levels of $\theta = 1$ the information function decays and the standard
427	error increases.
428	INSERT FIGURE 4

- 429
- 430

#### 431 **Discussion**

432 In the present study, we provided a psychometric analysis of a new measurement tool, the

433 CDDS, which is based on previous literature about coping and on findings from a preliminary

qualitative study. After the semi-structured interviews and the inter-rater agreement process, aninitial 48-item scale was obtained.

The results of the three quantitative studies enabled us to validate a 17-item scale (see Table 2) whose psychometric characteristics were globally satisfactory. The internal structure of the CDDS was analyzed via EFA, which revealed that a four-factor solution provided the best fit for this instrument. CFA confirmed this structure. It would be desirable for future research to test the factorial invariance across different groups of disability. With respect to the marker items and their content definition, we named the factors positive thinking, social sensitization and support, adaptation, and avoidance.

443 Although some of these factors, or similar ones (such as avoidance, cognitive coping and reinterpreting), emerged in prior research, this structure does not match previous results on 444 different groups and instruments<sup>1</sup> (Harland & Georgieff, 2010; Kara & Açıkel, 2011; Regier & 445 446 Parmalee, 2015; Skinner et al., 2003). This lack of concordance can be consistent with Skinner et al.'s (2003) claim that coping is an organizational construct used to encompass the many 447 different strategies people use to deal with stressful events. Since coping strategies may vary by 448 449 context, the studies carried out to develop and validate the CDDS may have identified the coping strategies that are relevant for facing the difficulties that arise from having a disability, 450 at least in the Spanish context. To our knowledge, this is the first coping scale that is 451 specifically designed for people with disabilities, and we think it can be highly useful for both 452 453 research and applied purposes.

<sup>&</sup>lt;sup>1</sup> Harland and Georgieff (in a sample of chronic back pain patients) found four factors: Catastrophizing, Diversion, Reinterpreting, and Cognitive coping. No equivalent to the first two was found in the present study, while the two latter are similar to our Positive thinking factor.

In their interviews with people with physical disabilities, Kara and Açıkel found three factors: Problem solving, Avoidance, and Seeking social support. This is the most similar to our structure, which includes Adaptation, Avoidance and Social sensitization and support.

Perhaps the most different is the structure found by Regier and Parmalee in a sample of older adults with osteoarthritis. They found five factors: Stoicism, Refocusing, Problem-Solving, Wishful Thinking, and Emotion-Focused. Only Refocusing and Problem-Solving are similar to factors found in the CDDS: Positive thinking and Adaptation.

Analyses regarding sociodemographic variables reveal significant differences in the four 454 455 factors. The greater use of approach strategies reported by blind participants might be specific 456 to the Spanish context, which is characterized as a traditionally well-organized community (ONCE, the National Organization of Blind Persons of Spain is among the oldest and most 457 recognized NGOs in Spain; it provides blind people with occupation, education, and resources). 458 459 In this sense, activism is associated with a greater use of approach strategies and a lower use of 460 avoidance strategies. Similarly, educational level might provide personal resources that help individuals cope with daily difficulties (and might also be associated with a higher economic 461 462 level, which was not controlled for), which may also be the reason that people with disabilities 463 at birth seem to use more positive coping strategies.

464 The results reflect significant differences in coping strategies between men and women. Women scored higher in social sensitization, support and adaptation. This result was similar to 465 findings of previous studies (Guszkowska, Zagórska-Pachucka, Kuk & Skwarek, 2016; Rose & 466 467 Rudolph, 2006; Tamres, Janicki, & Helgeson, 2002). Additionally, in a Spanish context, two studies that focused on gender differences found that women focused on emotional and 468 support-seeking strategies (Mataud, 2004; Meléndez, Mayordomo, Sancho, & Tomás, 2012). 469 These results have been regarded by different authors as the effect of socialization (Gattino, 470 471 Rollero, & De Piiccoli, 2015; Mataud, 2004.; Meléndez et al., 2012; Ptacek, Smith, & Zanas, 472 1992). Traditionally, women are taught to use more passive and emotionally and socially 473 focused behavior. Future studies on the role of gender in coping strategies should expand on 474 this topic.

The present study provides evidence that the CDDS has a high internal consistency and is able to distinguish between respondents who perceive themselves with more or less level of coping strategies, with good item discrimination both from the TCT and from the TRI perspective.

The scale is useful for measuring low and intermediate levels of the general trait that the CDDSscale measures, that is, the ability to use coping strategies to face the challenges posed by their

disability. So, if estimates are needed outside this range, then more items with thresholds above
1 are needing to measure the high levels of coping strategies. In sum, the findings of reliability
and validity indicate that CDDS is a reliable and useful scale to measure the coping strategies
used by people with disabilities to face daily difficulties.

The main limitation of this scale is that it was developed based on the responses of people with visual, hearing and physical disabilities. Thus, before it can be used with people with different types of disabilities (such as intellectual disability), it should be validated for those populations. Moreover, this scale was developed and tested in the Spanish context. Thus, the scale, and its factor structure, should be tested in samples from other cultures in order to ensure its applicability in different context.

490 Correlations found between each factor and other scales point to the relevance of the scale for 491 both theory and intervention. The case of avoidance is perhaps most interesting, as the use of 492 this strategy is associated with lower quality-of-life scores, higher levels of internalized stigma, 493 and lower levels of activism, perceived social support, self-esteem and resilience. Confirming 494 these relationships and identifying people who most frequently use this type of coping might 495 allow the implementation of training interventions to promote more positive coping strategies 496 (which, as seen in Table 3, show almost the reverse relationships with the other variables).

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Sample	N	Mean age	SD	Men	Women	Physical disabilities	Hearing impairments	Visual impairments	Other disabilities	Inborn disability	Acquired disability
Qualitative study	27	40.66	11.32	11 (40.74%)	16 (59.26%)	10 (37.04%)	10 (37.04%)	10       7       0       14         7.04%)       (25.92%)       (51.85)		14 (51.85%)	13 (48.15%)
Pilot study	80	47.67	9.12	36 (45%)	33 (41.3%)	40 (50%)	13 (16.3%)	9 (11.2%)	18 (23.6%)	18 (22.5%)	54 (67.5%)
EFA and CFA sample	590	44.76	12.78	269 (45.6%)	316 (53.6%)	354 (60%)	106 (18%)	96 (16.3%)	34 (5.8%)	207 (35.1%)	378 (64.1%)

# Table 1Sociodemographic and clinical data of the three samples

*Notes.* Except for the mean and *SD* of age, all data are frequencies. Some participants did not indicate their gender or the origin of their disabilities.

# Table 2

# <u>Results of the Exploratory Factor Analysis and descriptive statistics of the items</u> Items

	F1	F2	F3	F4	Lambda $(\lambda)$	Mean	Median	Kurtosis	Skewness
1. I try to look at the bright side of what happens to me and enjoy life	.59				.81	3.03	3.00	-0.50	-0.32
4. I keep my sense of humor	.72				.78	2.88	3.00	-0.73	-0.32
7. I value myself for what I have achieved	.59	.34			.70	3.08	3.00	-0.47	-0.44
8. Whenever I encounter a difficulty, I tell myself "come on, you can do it"	.59	.43			.80	3.11	3.00	-0.50	-0.42
13. Making fun of myself helps me go on	.60	.30			.67	2.77	3.00	-1.01	-0.21
16. I try to be as positive as possible	.57	.30			.91	3.13	3.00	-0.39	-0.51
2. I try to solve people's doubts about my disability		.79			.73	2.95	3.00	-0.61	-0.39
9. I speak publically about my disability to raise public awareness about the		.58			.66	2.36	2.00	-1.09	0.17
10. I ask other people for help when I need it		.30			.51	2.78	3.00	-0.85	-0.01
14. I try to make people see that people with disabilities are like everybody else.		.49			.86	3.17	3.00	-0.23	-0.80
19. I join other people to defend the rights of people with disabilities		.67			.58	2.24	2.00	-1.12	0.34
3. I organize my daily life in order to cope with the limitations arising from my disability in the best possible way			.61		.80	3.11	3.00	0.181	-0.51
15. I plan my daily activities			.73		.55	2.95	3.00	-0.58	-0.49
17. If I need any kind of resource, I look for it			.40		.79	3.20	3.00	-0.20	-0.50
5. I avoid attending social gatherings to spare myself difficulties				.79	.72	2.08	2	-0.79	0.46
12. I avoid situations which are difficult for me				.54	.60	2.64	3.00	-0.59	-0.12
20. I avoid interacting with other people				.69	.83	1.69	1.00	0.29	1.07

*Note*. Values < .30 are not shown in the table.

#### Table 3.

# Scores, covariances and correlations of the CDDS factors

Covariances ( $\phi$ )Pearson's Correlations ( $r$ )													
	$M\pm SD$	РТ	SSS	AD	AV	Quality	Perceived	Internalized	Group	Activism	Social	Self-	Resilience
						of Life	Discrimination	stigma	identification		Support	esteem	
PT	3.00±.64	1	.69	.80	65	.57**	09*	44**	.12**	.21**	.37**	.68**	.73**
SSS	2.71±.67		1	.72	58.	.37**	.08	22**	.38**	.39**	.44**	.36**	.38**
AD	3.08±.61			1	60	.48**	.00	26**	.19**	.23**	.31**	.50**	.49**
AV	2.14±.68				1	51**	.31**	.55**	.02	13**	40**	52**	42**

*Note*. The scale has a range of four points. PT = Positive Thinking; SSS = Social sensitization and support; AD = Adaptation; AV = Avoidance.

\* *p* < .05, \*\* *p* < .01

#### Table 4

#### Fit indexes of the proposed model and the alternative models

	RMSEA [90% CI]	SRMR	GFI	AGFI	CFI	NFI	NNFI	IFI	CAIC	Satorra-	df	$\Delta \chi^2 (\Delta df)^1$
										Bentler $\chi^2$		
Model 1. (four factors)	.056 [044 ; .068]	.061	.98	.96	.98	.97	.98	.98	468.17	206.18 ( <i>p</i> < .001)	113	
Model 2.												
(two factors grouping the three approach factors)	.085 [0.75 ; 0.96]	.078	.97	.96	.96	.94	.96	.96	569.75	339.99 ( <i>p</i> < .001)	118	133.81 (5) ( <i>p</i> < .001)
Model 3. (three factors grouping PT+AD in factor 1)	.061 [049 ; 0.072]	.066	.98	.97	.98	.96	.98	.98	469.60	226.71 ( <i>p</i> < .001)	116	20.53 (3) ( <i>p</i> < .001)
Model 4.	002 [ 092 - 10]	096	06	05	06	0.4	05	06	604 71	381.52		175.34 (6)
(one single factor)	.092 [.082 ; .10]	.086	.96	.95	.96	.94	.95	.96	604.71	( <i>p</i> < .001)	119	( <i>p</i> < .001)

<sup>1</sup>Change in chi-square between the proposed model (Model 1) and the rest of the proposed alternative models.

Model 1. The four-factor proposed model. Model 2. Two factors (Approach (PT+SSS+AD) and Avoidance). Model 3: Three factors. We have tested the model grouping PT and AD in factor 1 because the high covariance between the two factors Model 4: one single factor.

#### Table 5

Factors	Cronbach's	Average	Jöreskog
	Alfa	Variance	rho
		Extracted	(Omega)
		(AVE)	-
Positive thinking (PT)	.86	.60	.90
Social sensitization and support (SSS)	.75	.50	.80
Adaptation (AD)	.69	.61	.76
Avoidance (AV)	.68	.61	.76

Fully standardized solution of the model, construct and reliability statistics

Items	а	$b_1$	$b_2$	$b_3$	$S-\chi^2$	р
Item 1	2.49 (.20)	-2.40 (.16)	-0.97 (.07)	0.60 (.08)	59.71	.1402
Item 2	1.28 (.12)	-2.68 (.23)	-0.96 (.10)	0.84 (.12)	87.94	.0841
Item 3	1.69 (.15)	-2.96 (.24)	-1.40 (.10)	0.73 (.10)	77.31	.0379
Item 4	2.05 (.16)	-1.98 (.13)	-0.63 (.07)	0.71 (.09)	84.50	.0202
Item 5	0.90 (.10)	-2.91 (.31)	-0.99 (.13)	0.87 (.15)	86.07	.3571
Item 7	1.74 (.15)	-2.81 (.22)	-1.07 (.09)	0.56 (.09)	91.03	.0016
Item 8	2.36 (.20)	-2.64 (.19)	-1.05 (.07)	0.51 (.08)	67.45	.0268
Item 9	1.04 (.10)	-1.49 (.15)	0.17 (.10)	1.68 (.19)	105.01	.0376
Item 10	0.83 (.10)	-4.20 (.49)	-0.69 (.13)	1.64 (.22)	106.71	.0077
Item 12	0.62 (.09)	-3.00 (.43)	0.42 (.16)	3.87 (.56)	111.42	.0288
Item 13	1.61 (.13)	-1.85 (.14)	-0.42 (.07)	0.83 (.10)	79.91	.2190
Item 14	1.58 (.14)	-2.39 (.19)	-1.18 (.10)	0.20 (.08)	77.21	.1424
Item 15	1.12 (.11)	-2.77 (.26)	-1.05 (.12)	0.84 (.13)	97.06	.0373
Item 16	3.10 (.29)	-2.25 (.14)	-1.00 (.06)	0.38 (.07)	68.59	.0078
Item 17	2.01 (.18)	-3.07 (.27)	-1.38 (.09)	.45 (.08)	65.91	.0439
Item 19	0.86 (.10)	-1.15 (.15)	0.52 (.13)	2.07 (.25)	110.22	.0245
Item 20	1.05 (.12)	-3.46 (.37)	-1.90 (.19)	-0.23 (.10)	125.93	.0001

SGR model item parameter estimates and Item-Fit Statistics for the Coping with Disability Difficulties Scale (CDDS)

*Note.* a = item slope (discrimination) parameter; b = item threshold (location) parameter;  $S \cdot \chi^2 =$  item fit statistic; p = p value associated with item-fit statistic. Values in parenthesis are item parameter standard error estimate.

Steps in the research	Procedure	Participants
Step 1		
Literature review	Identification of the initial contents and domains.	Research team.
Step 2		
Qualitative study	Discourse analysis for the identification of the initial contents and domains.	N = 27 people with disabilities (physical, visual and hearing).
Step 3	I	
Development of initial item pool	Construct definition and item development	Research team.
(110 items)	Inter-rater agreement process.	
	Cognitive debriefing.	N = 8 people with disabilities of different types and severity.
Step 4		
Pilot study (46 items)	Analysis of the items' psychometric properties.	N = 80 people with disabilities (physical, visual and hearing).
Step 5		
Validation studies (20 items)	Construct validity analysis for the instrument.	N = 52 <u>Q</u> <sup>1</sup> people with disabilities (physical, visual and hearing).
	Cross-validation of factorial structure and convergent validity analysis of the instrument.	

Figure 1. Steps of the process of development of the Coping with Disability Difficulties Scale.



*Figure 3.* Option Response Functions (ORF) for items of the CDDS scale fit by the GRS model.



*Figure 4*. Total information curve for the CDDS scale (function test information: continuous line; standard error of measurement: dotted line).